

THE ROYAL  
SWEDISH  
ACADEMY OF  
SCIENCES



**INSTITUT  
MITTAG-LEFFLER**

Auravägen 17, SE-182 60 Djursholm, Sweden  
Tel. +46 8 622 05 60 Fax. +46 8 622 05 89  
info@mittag-leffler.se www.mittag-leffler.se

**On the distributions of the sup and inf of the  
classical risk process with exponential claim**

J. León and J. Villa

REPORT No. 8, 2007/2008, fall

ISSN 1103-467X

ISRN IML-R- -8-07/08- -SE+fall

# ON THE DISTRIBUTIONS OF THE SUP AND INF OF THE CLASSICAL RISK PROCESS WITH EXPONENTIAL CLAIM\*

JORGE A. LEÓN AND JOSÉ VILLA

ABSTRACT. The purpose of this article is to use the double Laplace transform of the occupation measure of the classical risk process  $X$  with exponential claim to deduce the distributions of the random variables  $\sup\{X_s : s \leq t\}$  and  $\inf\{X_s : s \leq t\}$ , for every  $t > 0$ . As a consequence, we also get the distributions of the time to ruin in finite time and the first passage of a given level.

## 1. INTRODUCTION

In this paper, we deal with the classical risk process with exponential claim defined on a complete probability space  $(\Omega, \mathcal{F}, P)$ . More precisely, let

$$(1.1) \quad X_t = x_0 + ct - \sum_{k=1}^{N_t} R_k, \quad t \geq 0.$$

Here  $x_0 \geq 0$  is the initial capital,  $c > 0$  is the premium income per unit of time,  $N = \{N_t, t \geq 0\}$  is an homogeneous Poisson process with rate  $\lambda$  and  $\{R_k, k = 1, 2, \dots\}$  is a sequence of i.i.d. random variables independent of  $N$ . Henceforth we suppose that  $R_1$  has exponential distribution with mean  $1/r$ .

Our goal is to calculate explicit expressions for the distributions of the random variables  $\sup\{X_s : s \leq t\}$  and  $\inf\{X_s : s \leq t\}$ ,  $t > 0$ . Toward this end, we apply the complex inversion theorem of the Laplace transform (or Lerch's theorem) to the double Laplace transforms of some occupation measures of  $X$  (see Section 2 below). As a consequence, we are also able to give the distribution of the first passage of certain level  $x \in \mathbb{R}$  of the process  $X$ . It means, the distributions of

$$S_x = \inf\{t > 0 : X_t = x\} \quad \text{and} \quad T_x = \inf\{t > 0 : X_t < x\},$$

because the right-continuity of the process  $X$  yields

$$\{S_x < t\} = \{\sup_{s \leq t} X_s > x\}, \quad x > x_0$$

and

$$\{T_x \leq t\} = \{\inf_{s \leq t} X_s < x\}, \quad x < x_0.$$

---

2000 *Mathematics Subject Classification.* Primary 60K30; Secondary 60K99.

*Key words and phrases.* Classical risk process, Laplace transform, ruin probability, time to ruin.

\*Partially supported by the CONACyT grant 45684-F and by PIM 08-2 of UAA.

The study of the distribution of the time to ruin  $T_0$  in finite time (i.e., for the case  $x = 0$ ) is extensive in the literature on risk theory due to its applications in business activities. For instance, numerical procedures have been utilized by several authors in the analysis of  $T_0$  (see, for example, Dickson and Waters [6] or Seal [12]). This numerical approximations have been improved in several works by deriving expressions for the mentioned distribution (see Asmussen [1, 2], Dickson [4], Dickson et al. [5, 7], Drekić and Willmot [9], and Ignatov and Kaishev [11], among others). In particular, the method used in [5], [7] and [9] is based on the complex inversion theorem, as we does here.

For  $x \in \mathbb{R}$ , the process  $S_x$  have been analyzed by doss Reis [8] and Gerber [10] via a martingale method.

The paper is organized as follows. In Section 2 we relate the Laplace transforms of the functions

$$(1.2) \quad P\{\sup_{s \leq \cdot} X_s \leq x\} \quad \text{and} \quad P\{\inf_{s \leq \cdot} X_s \leq x\}$$

to the double Laplace transform of some occupation measures of  $X$ . Then, in Section 3 we use the complex inversion theorem to calculate the two probabilities in (1.2).

## 2. OCCUPATION MEASURES

Now we are interested in the Laplace transforms of the occupation measures

$$Y_x(t) = \int_0^t 1_{(x, +\infty)}(X_s) ds \quad \text{and} \quad Y^x(t) = \int_0^t 1_{(-\infty, x)}(X_s) ds,$$

with  $x \in \mathbb{R}$  and  $t > 0$ . So, we assume that the reader is familiar with the elementary properties of the Laplace transform as they are presented, for example, in Spiegel [13].

Throughout, the Laplace transform of a measurable function  $h : [0, \infty) \rightarrow \mathbb{R}$  is denoted by  $\mathcal{L}(h)$ . That is,

$$\mathcal{L}(h)(s) = \int_0^\infty e^{-st} h(t) dt,$$

for  $s \in \mathbb{R}$  such that this integral is convergent.

The relation between the occupation measures  $Y_x$  and  $Y^x$ , and the probabilities in (1.2) is given by the following result.

**Proposition 1.** *Let  $X$  be the classical risk process defined in (1.1) and  $x \in \mathbb{R}$ . Then for each  $t > 0$ ,*

$$\{Y_x(t) = 0\} = \left\{ \sup_{s \leq t} X_s \leq x \right\} \quad \text{and} \quad \{Y^x(t) = 0\} = \left\{ \inf_{s \leq t} X_s \geq x \right\}.$$

*Proof.* We first observe that  $\{Y_x(t) = 0\} \supset \{\sup_{s \leq t} X_s \leq x\}$  is trivial. Now we see the reverse inclusion. Let  $\omega \in \Omega$  be such that

$$(2.1) \quad \int_0^t 1_{(x, +\infty)}(X_s(\omega)) ds = 0.$$

If there is some  $s_0 \in (0, t)$  such that  $X_{s_0}(\omega) > x$ , then, by the right-continuity of  $X$ , there exists a non-empty open interval  $I_{s_0} \subset (0, t)$  such that  $s_0 \in I_{s_0}$  and  $X_s(\omega) > x$ , for all  $s \in I_{s_0}$ . Consequently,

$$\int_0^t 1_{(x, +\infty)}(X_s(\omega)) ds \geq \int_{I_{s_0}} 1_{(x, +\infty)}(X_s(\omega)) ds = |I_{s_0}| > 0,$$

where  $|I_{s_0}|$  is the length of  $I_{s_0}$ . But this is a contradiction to (2.1). Therefore  $X_s(\omega) \leq x$  for all  $s \leq t$ , which implies that  $\omega$  also belongs to  $\{\sup_{s \leq t} X_s \leq x\}$ .

We proceed similarly for the remainder of the proof.  $\square$

In order to express the double Laplace transform of  $Y^x$  and  $Y_x$ , we need to introduce the following notation. Let  $\tilde{n}$  be a positive real number. The positive and negative roots of the quadratic equation

$$cs^2 + (rc - \lambda - \tilde{n})s - \tilde{n}r = 0$$

are denoted by  $v_n^+$  and  $v_n^-$ , respectively.

**Proposition 2.** *Let  $s$  and  $\alpha$  be two positive real numbers. Then*

$$\int_0^\infty e^{-st} E \left[ e^{-\alpha Y_x(t)} \right] dt = \begin{cases} \frac{1}{s+\alpha} \left\{ 1 + \frac{\alpha(\frac{1}{c} - \frac{v_s^+}{s})}{v_{s+\alpha}^+ - v_s^+ - \frac{\alpha}{c}} e^{(x_0-x)v_{s+\alpha}^-} \right\}, & x_0 \geq x, \\ \frac{1}{s} + \frac{\frac{\alpha sc}{s+\alpha} (\frac{1}{c} - \frac{v_s^+}{s}) v_{s+\alpha}^- - \alpha (v_{s+\alpha}^+ - v_s^+ - \frac{\alpha}{c})}{sc(v_s^+ + \frac{\alpha}{c})(v_{s+\alpha}^- - v_s^+ - \frac{\alpha}{c})} e^{(x-x_0)v_s^+}, & x_0 < x, \end{cases}$$

and

$$\int_0^\infty e^{-st} E \left[ e^{-\alpha Y^x(t)} \right] dt = \begin{cases} \frac{1}{s+\alpha} \left\{ 1 + \frac{\alpha(\frac{1}{c} - \frac{v_s^-}{s})}{v_{s+\alpha}^+ - v_s^- - \frac{\alpha}{c}} e^{(x_0-x)v_{s+\alpha}^+} \right\}, & x_0 \leq x, \\ \frac{1}{s} + \frac{\frac{\alpha sc}{s+\alpha} (\frac{1}{c} - \frac{v_s^-}{s}) v_{s+\alpha}^+ - \alpha (v_{s+\alpha}^+ - v_s^- - \frac{\alpha}{c})}{sc(v_s^- + \frac{\alpha}{c})(v_{s+\alpha}^+ - v_s^- - \frac{\alpha}{c})} e^{(x-x_0)v_s^-}, & x_0 \geq x. \end{cases}$$

*Proof.* The proposition is an immediate consequence of Chiu and Yin [3] (Corollary 4.1).  $\square$

The following result is a consequence of Proposition 2 and it will be used in Section 3.

**Proposition 3.** *Let  $X$  be the classical risk process given by (1.1). Then, for every  $s > 0$ , we have*

$$(2.2) \quad \mathcal{L}(P(Y_x(\cdot) = 0))(s) = \begin{cases} 0, & x_0 \geq x, \\ \frac{1}{s} - \frac{e^{-(x-x_0)v_s^+}}{s}, & x_0 < x, \end{cases}$$

and

$$(2.3) \quad \mathcal{L}(P(Y^x(\cdot) = 0))(s) = \begin{cases} 0, & x_0 \leq x, \\ \frac{1}{s} - \frac{\lambda e^{-(x-x_0)v_s^-}}{cs(v_s^+ + r)}, & x_0 > x. \end{cases}$$

*Proof.* We first deal with equality (2.2).

By the dominated convergence theorem we can write

$$\begin{aligned}
& \lim_{\alpha \rightarrow \infty} \int_0^\infty e^{-st} E \left[ e^{-\alpha Y_x(t)} \right] dt \\
&= \lim_{\alpha \rightarrow \infty} \int_0^\infty e^{-st} \left( \int_{\{Y_x(t)=0\}} + \int_{\{Y_x(t)>0\}} \right) e^{-\alpha Y_x(t)} dP dt \\
&= \int_0^\infty e^{-st} P(Y_x(t) = 0) dt + \int_0^\infty e^{-st} \int_{\{Y_x(t)>0\}} \left( \lim_{\alpha \rightarrow \infty} e^{-\alpha Y_x(t)} \right) dP dt \\
&= \int_0^\infty e^{-st} P(Y_x(t) = 0) dt.
\end{aligned}$$

Therefore, from Proposition 2 we get

$$\begin{aligned}
(2.4) \quad & \int_0^\infty e^{-st} P(Y_x(t) = 0) dt \\
&= \lim_{\alpha \rightarrow \infty} \begin{cases} \frac{1}{s+\alpha} \left\{ 1 + \frac{\alpha \left( \frac{1}{c} - \frac{v_s^+}{s} \right)}{v_{s+\alpha}^- - v_s^+ - \frac{\alpha}{c}} e^{(x_0-x)v_{s+\alpha}^-} \right\}, & x_0 \geq x, \\ \frac{1}{s} + \frac{\frac{\alpha s c}{s+\alpha} \left( \frac{1}{c} - \frac{v_s^+}{s} \right) v_{s+\alpha}^- - \alpha \left( v_{s+\alpha}^- - v_s^+ - \frac{\alpha}{c} \right)}{s c \left( v_s^+ + \frac{\alpha}{c} \right) \left( v_{s+\alpha}^- - v_s^+ - \frac{\alpha}{c} \right) e^{(x-x_0)v_s^+}}, & x_0 < x. \end{cases}
\end{aligned}$$

Note that the definition of  $v_{s+\alpha}^-$  implies that  $\lim_{\alpha \rightarrow \infty} (v_{s+\alpha}^-/\alpha) = 0$ , which leads us to

$$\lim_{\alpha \rightarrow \infty} \left( \alpha \left( v_{s+\alpha}^- - v_s^+ - \frac{\alpha}{c} \right)^{-1} \right) = -c.$$

Hence, the fact that  $v_{s+\alpha}^- < 0$ , for all  $\alpha > 0$ , together with (2.4), yields that equality (2.2) is true for  $x_0 \geq x$ .

On the other hand, for  $x_0 < x$ ,

$$\begin{aligned}
& \lim_{\alpha \rightarrow \infty} \left( \frac{1}{s} + \frac{\frac{\alpha s c}{s+\alpha} \left( \frac{1}{c} - \frac{v_s^+}{s} \right) v_{s+\alpha}^- - \alpha \left( v_{s+\alpha}^- - v_s^+ - \frac{\alpha}{c} \right)}{s c \left( v_s^+ + \frac{\alpha}{c} \right) \left( v_{s+\alpha}^- - v_s^+ - \frac{\alpha}{c} \right) e^{(x-x_0)v_s^+}} \right) \\
&= \lim_{\alpha \rightarrow \infty} \left( \frac{1}{s} + \frac{s c \left( \frac{1}{c} - \frac{v_s^+}{s} \right) \frac{v_{s+\alpha}^-}{s+\alpha} - \left( v_{s+\alpha}^- - v_s^+ - \frac{\alpha}{c} \right)}{s c \left( \frac{v_s^+}{\alpha} + \frac{1}{c} \right) \left( v_{s+\alpha}^- - v_s^+ - \frac{\alpha}{c} \right)} e^{-(x-x_0)v_s^+} \right) \\
&= \frac{1}{s} - \frac{e^{-(x-x_0)v_s^+}}{s}.
\end{aligned}$$

Thus (2.2) holds.

Finally, in order to see that (2.3) is satisfied, we only need to observe that

$$\lim_{\alpha \rightarrow \infty} (v_{s+\alpha}^+/\alpha) = \frac{1}{c} \quad \text{and} \quad \lim_{\alpha \rightarrow \infty} \left( v_{s+\alpha}^+ - v_s^- - \frac{\alpha}{c} \right) = v_s^+ + r,$$

and proceed as in the beginning of this proof.  $\square$

3. THE DISTRIBUTIONS OF THE SUP AND INF OF  $X$  WITHIN FINITE TIME

The purpose of this section is to apply Lerch's theorem to the Laplace transforms obtained in Proposition 3 to calculate the distributions of the sup and inf of  $X$  (see Proposition 1).

In order to state the main result of this paper, we need to introduce the following notation.

Note first that

$$\begin{aligned} (s + \lambda - rc)^2 + 4crs &= (s + \lambda + rc)^2 - 4\lambda rc \\ &= (s + \lambda + rc - 2\sqrt{\lambda rc})(s + \lambda + rc + 2\sqrt{\lambda rc}) \\ &= (s - r_1)(s - r_2), \end{aligned}$$

with  $r_1 = 2\sqrt{\lambda rc} - \lambda - rc$  and  $r_2 = -2\sqrt{\lambda rc} - \lambda - rc$ . Hence  $r_2 < r_1 < 0$ ,

$$(3.1) \quad v_s^+ = \frac{s + \lambda - rc + \sqrt{(s - r_1)(s - r_2)}}{2c},$$

and

$$(3.2) \quad v_s^- = \frac{s + \lambda - rc - \sqrt{(s - r_1)(s - r_2)}}{2c}.$$

For sake of simplicity, let us utilize the conventions  $a = \frac{x - x_0}{2c}$ ,  $b = \lambda + rc$  and

$$\chi(z) = \frac{\arg(z - r_1) + \arg(z - r_2)}{2}, \quad z \in \mathbb{C}.$$

**Theorem 1.** *Let  $X$  be the classical risk process (1.1) and  $t > 0$ . Then we have*

$$(3.3) \quad P\left(\sup_{s \leq t} X_s > x\right) = e^{-a(\lambda - rc)} \left( e^{-a\sqrt{r_1 r_2}} + \frac{1}{\pi} \int_{r_2}^{r_1} \frac{e^{(t-a)u} \sin(a|u - r_1|^{1/2}|u - r_2|^{1/2})}{u} du \right),$$

for every  $x \in (x_0, x_0 + ct)$ , and

$$(3.4) \quad P\left(\inf_{s \leq t} X_s < x\right) = 2\lambda e^{-a(\lambda - rc)} \left( \frac{e^{a\sqrt{r_1 r_2}}}{b + \sqrt{r_1 r_2}} + \frac{1}{\pi} \operatorname{Im} \left( \int_{r_2}^{r_1} \frac{e^{(t-a)u - ai|u - r_1|^{1/2}|u - r_2|^{1/2}}}{u(u + b - i|u - r_1|^{1/2}|u - r_2|^{1/2})} du \right) \right),$$

for every  $x < x_0$ .

**Remarks.**

i) Note that

$$P\left(\sup_{s \leq t} X_s > x\right) = 0 \text{ for } x \geq x_0 + ct, \quad \text{and} \quad P\left(\sup_{s \leq t} X_s > x\right) = 1 \text{ for } x < x_0.$$

ii) Similarly we have

$$P\left(\inf_{s \leq t} X_s < x\right) = 1 \quad \text{for } x > x_0.$$

In the following two subsections, we separate the proofs of (3.3) and (3.4) for the convenience of the reader because both of them are long and tedious.

**3.1. Proof of equality (3.3).** Let

$$h_M(s) = \frac{e^{-as-a\sqrt{(s-r_1)(s-r_2)}}}{s}, \quad s \in \mathbb{C} \setminus [r_2, r_1],$$

with  $\sqrt{(s-r_1)(s-r_2)} = |(s-r_1)(s-r_2)|^{1/2} \exp(i\chi(s))$ . Then, by (2.2), (3.1), the inverse theorem of the Laplace transform (see for example [13]) and the fact that  $h_M$  is an analytic function on  $\mathbb{C} \setminus [r_2, r_1]$ , we have, for  $\sigma$  large enough,

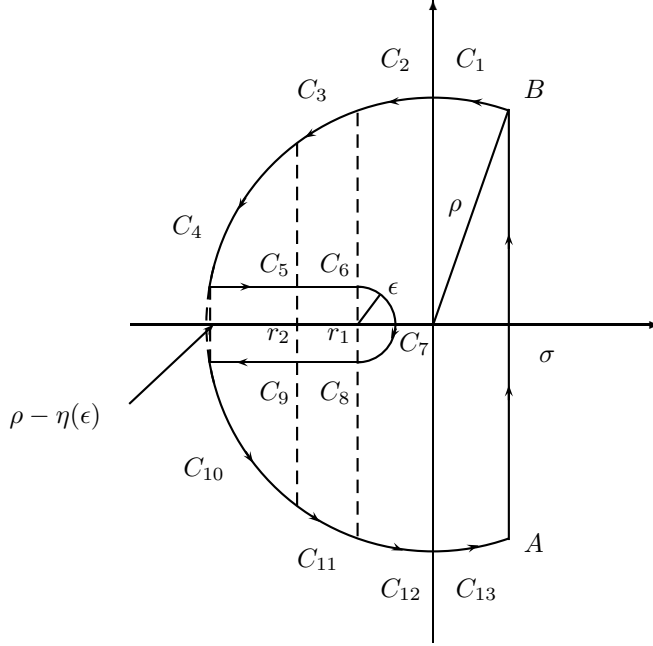
$$\begin{aligned} P(Y_x(t) = 0) &= \frac{1}{2\pi i} \int_{\sigma-i\infty}^{\sigma+i\infty} e^{ts} \left( \frac{1}{s} - \frac{e^{-(x-x_0)v_s^+}}{s} \right) ds \\ &= 1 - \frac{1}{2\pi i} \int_{\sigma-i\infty}^{\sigma+i\infty} e^{ts} \frac{e^{a(rc-\lambda-s-\sqrt{(s-r_1)(s-r_2)})}}{s} ds \\ &= 1 - \frac{e^{a(rc-\lambda)}}{2\pi i} \int_{\sigma-i\infty}^{\sigma+i\infty} e^{ts} h_M(s) ds. \end{aligned}$$

Notice that we can use the inverse theorem because it is not difficult to see that  $t \mapsto P(X_t = x) = 0$  is continuous, which follows from the fact that  $P(X_t = x) = 0$ .

Now observe that 0 is a pole of order one and  $r_1, r_2$  are branch points of  $h_M$ . Therefore, by the residue theorem (see [13]),

$$(3.5) \quad \int_{\sigma-i\infty}^{\sigma+i\infty} e^{ts} h_M(s) ds = 2\pi i e^{-a\sqrt{r_1 r_2}} - \lim_{\rho \rightarrow \infty} \lim_{\varepsilon \rightarrow 0} \int_{C(\rho, \varepsilon)} e^{ts} h_M(s) ds.$$

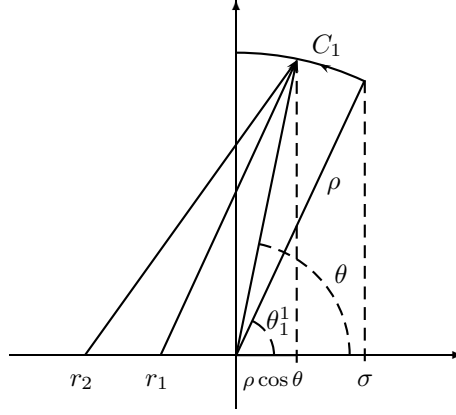
Here  $C(\rho, \varepsilon) = C_1 \cup \dots \cup C_{13}$  is the following contour of integration:



Note that we only need to analyze the integral in the right-hand side of (3.5) on each arc  $C_j$ ,  $j \in \{1, \dots, 13\}$ , in order to finish the proof. To do so, now we divide the proof in several steps.

**Step 1.** We begin our study on the arcs  $C_1(\rho)$  and  $C_{13}(\rho)$ .

For  $C_1(\rho)$  we take the parametrization  $s = \rho e^{i\theta}$ ,  $\theta_1^1 \leq \theta \leq \pi/2$ , where  $\rho$  and  $\theta_1^1$  are indicated in the following figure:



In this case, it is easy to see that

$$0 \leq \arg(\rho e^{i\theta} - r_1) < \frac{\pi}{2} \quad \text{and} \quad 0 \leq \arg(\rho e^{i\theta} - r_2) < \frac{\pi}{2},$$

which give  $0 \leq \chi(\rho e^{i\theta}) < \pi/2$ . Since  $\cos \theta \geq 0$ , when  $\theta \in [0, \pi/2]$ , we have

$$\operatorname{Re}(\sqrt{(\rho e^{i\theta} - r_1)(\rho e^{i\theta} - r_2)}) = |\rho e^{i\theta} - r_1|^{1/2} |\rho e^{i\theta} - r_2|^{1/2} \cos(\chi(\rho e^{i\theta})) > 0.$$

Using this and the fact that  $a > 0$ , we can conclude

$$(3.6) \quad e^{-a \operatorname{Re}(\sqrt{(\rho e^{i\theta} - r_1)(\rho e^{i\theta} - r_2)})} \leq 1.$$

Now note that  $t > a$ , due to  $x_0 + ct > x$ , and that  $\rho \cos \theta < \sigma$ . Thus,

$$\begin{aligned} \left| \int_{C_1(\rho)} e^{ts} h_M(s) \right| &= \left| \int_{\theta_1^1}^{\pi/2} \frac{e^{(t-a)\rho e^{i\theta} - a \sqrt{(\rho e^{i\theta} - r_1)(\rho e^{i\theta} - r_2)}}}{\rho e^{i\theta}} \rho i e^{i\theta} d\theta \right| \\ &\leq \int_{\theta_1^1}^{\pi/2} e^{(t-a)\rho \cos \theta - a \operatorname{Re}(\sqrt{(\rho e^{i\theta} - r_1)(\rho e^{i\theta} - r_2)})} d\theta \\ &\leq \int_{\theta_1^1}^{\pi/2} e^{(t-a)\sigma} d\theta \\ &= e^{(t-a)\sigma} \sin^{-1} \left( \frac{\sigma}{\rho} \right) \leq e^{(t-a)\sigma} \left( \frac{\sigma}{\rho} \right) \frac{\pi}{2}. \end{aligned}$$

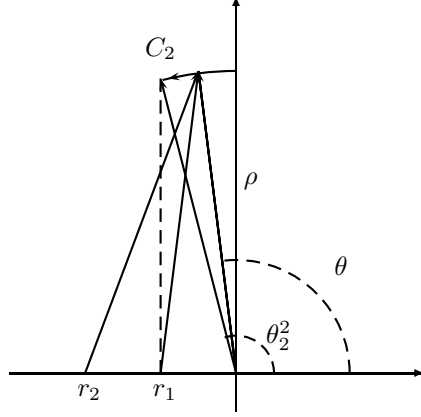
Hence  $\lim_{\rho \rightarrow \infty} \int_{C_1(\rho)} e^{ts} h_M(s) ds = 0$ .

We can proceed in the same way to see that  $\lim_{\rho \rightarrow \infty} \int_{C_{13}(\rho)} e^{ts} h_M(s) ds$  is also equal to zero.

Note that an important point in this analysis is the fact that  $t - a > 0$  and  $-a < 0$ . This will be also important in the remaining of this proof.

**Step 2.** Now we consider the integral over  $C_2(\rho)$  and  $C_{12}(\rho)$ .

Over  $C_2(\rho)$ , we consider the parametrization  $s = \rho e^{i\theta}$ ,  $\pi/2 \leq \theta \leq \theta_2^2 < \pi$ , as it is illustrated in the next figure:



As in the previous case, the inequality (3.6) is still true. So, taking into account that  $\sin \theta \geq 2\theta/\pi$ , for  $\theta \in [0, \pi/2]$ , we conclude

$$\begin{aligned}
\left| \int_{C_2(\rho)} e^{ts} h_M(s) ds \right| &\leq \int_{\frac{\pi}{2}}^{\theta_2^2} e^{(t-a)\rho \cos \theta - a \operatorname{Re}(\sqrt{(\rho e^{i\theta} - r_1)(\rho e^{i\theta} - r_2)})} d\theta \\
&\leq \int_{\frac{\pi}{2}}^{\theta_2^2} e^{(t-a)\rho \cos \theta} d\theta \\
&= \int_0^{\theta_2^2 - \frac{\pi}{2}} e^{(t-a)\rho \cos(\theta + \frac{\pi}{2})} d\theta \\
&= \int_0^{\theta_2^2 - \frac{\pi}{2}} e^{-(t-a)\rho \sin \theta} d\theta \\
&\leq \int_0^{\theta_2^2 - \frac{\pi}{2}} e^{-(t-a)\rho(2\theta/\pi)} d\theta \leq \frac{\pi}{2\rho(t-a)}.
\end{aligned}$$

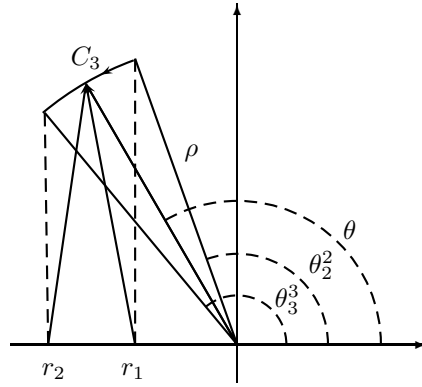
From which  $\lim_{\rho \rightarrow \infty} \int_{C_2(\rho)} e^{ts} h_M(s) ds = 0$ .

Similarly,  $\lim_{\rho \rightarrow \infty} \int_{C_{12}(\rho)} e^{ts} h_M(s) ds = 0$ .

**Step 3.** Here we will show that

$$\lim_{\rho \rightarrow \infty} \int_{C_3(\rho)} e^{ts} h_M(s) ds = 0 \quad \text{and} \quad \lim_{\rho \rightarrow \infty} \int_{C_{11}(\rho)} e^{ts} h_M(s) ds = 0.$$

On  $C_3(\rho)$ , we still use the parametrization  $s = \rho e^{i\theta}$ , with  $\frac{\pi}{2} < \theta_2^2 \leq \theta \leq \theta_3^3 < \pi$ :



Since  $x < x_0 + ct$ , then  $t > 2a$ . Therefore we can take  $\eta > 0$  such that

$$(3.7) \quad t > \left(1 + (1 + \eta)^{1/4}\right) a.$$

Moreover take  $\rho > 0$  such that

$$(3.8) \quad \eta \rho^2 > |r_2 - r_1|^2.$$

Notice that

$$\frac{\pi}{2} \leq \arg(\rho e^{i\theta} - r_1) < \theta \quad \text{and} \quad 0 \leq \arg(\rho e^{i\theta} - r_2) \leq \frac{\pi}{2}.$$

Hence

$$\frac{\pi}{4} \leq \chi(\rho e^{i\theta}) < \frac{\theta + \frac{\pi}{2}}{2} < \theta.$$

Since  $\cos$  is decreasing on  $[\pi/4, \pi]$ , we have

$$(3.9) \quad \operatorname{Re}(\sqrt{(\rho e^{i\theta} - r_1)(\rho e^{i\theta} - r_2)}) \geq |\rho e^{i\theta} - r_1|^{1/2} |\rho e^{i\theta} - r_2|^{1/2} \cos \theta.$$

On the other hand, the fact that  $r_1 \geq \rho \cos \theta$  implies

$$(3.10) \quad |\rho e^{i\theta} - r_1| \leq \rho.$$

And using (3.8), we get

$$(3.11) \quad \begin{aligned} |\rho e^{i\theta} - r_2| &\leq \sqrt{|r_1 - r_2|^2 + \rho^2 \sin^2(\theta)} \\ &\leq \sqrt{|r_1 - r_2|^2 + \rho^2} \\ &\leq (\eta + 1)^{1/2} \rho. \end{aligned}$$

Therefore (3.9), (3.10) and (3.11) yields

$$(3.12) \quad \begin{aligned} \operatorname{Re}(\sqrt{(\rho e^{i\theta} - r_1)(\rho e^{i\theta} - r_2)}) &\geq \rho^{1/2} (\eta + 1)^{1/4} \rho^{1/2} \cos \theta \\ &= (\eta + 1)^{1/4} \rho \cos \theta. \end{aligned}$$

From this and (3.7) we obtain

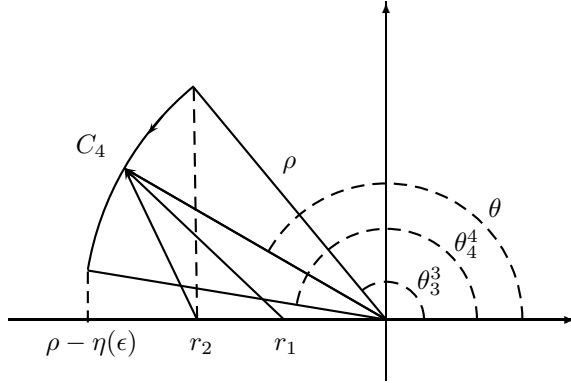
$$\begin{aligned}
\left| \int_{C_3(\rho)} e^{ts} h_M(s) ds \right| &\leq \int_{\theta_2^2}^{\theta_3^3} e^{(t-a)\rho \cos \theta - a \operatorname{Re}(\sqrt{(\rho e^{i\theta} - r_1)(\rho e^{i\theta} - r_2)})} d\theta \\
&\leq \int_{\theta_2^2}^{\theta_3^3} e^{(t-a)\rho \cos \theta - a(\eta+1)^{1/4} \rho \cos \theta} d\theta \\
&= \int_{\theta_2^2}^{\theta_3^3} e^{(t-a(1+(\eta+1)^{1/4}))\rho \cos \theta} d\theta \\
&= \int_{\theta_2^2 - \frac{\pi}{2}}^{\theta_3^3 - \frac{\pi}{2}} e^{(t-a(1+(\eta+1)^{1/4}))\rho \cos(\theta + \frac{\pi}{2})} d\theta \\
&= \int_{\theta_2^2 - \frac{\pi}{2}}^{\theta_3^3 - \frac{\pi}{2}} e^{-(t-a(1+(\eta+1)^{1/4}))\rho \sin \theta} d\theta \\
&\leq \int_{\theta_2^2 - \frac{\pi}{2}}^{\theta_3^3 - \frac{\pi}{2}} e^{-(t-a(1+(\eta+1)^{1/4}))\rho \frac{2\theta}{\pi}} d\theta \\
&\leq \frac{\pi}{2\rho(t-a(1+(\eta+1)^{1/4}))}.
\end{aligned}$$

This implies  $\lim_{\rho \rightarrow \infty} \int_{C_3(\rho)} e^{ts} h_M(s) ds = 0$ .

Proceeding as the beginning of this step, we can conclude that we also have  $\lim_{\rho \rightarrow \infty} \int_{C_{11}(\rho)} e^{ts} h_M(s) ds = 0$ .

**Step 4.** Now we deal with the arcs  $C_4(\rho, \varepsilon)$  and  $C_{10}(\rho, \varepsilon)$ .

Here we consider the same parametrization of previous steps. That is,  $s = \rho e^{i\theta}$ ,  $\pi/2 \leq \theta_3^3 \leq \theta \leq \theta_4^4 \leq \pi$ :



Notice that

$$\frac{\pi}{2} \leq \arg(\rho e^{i\theta} - r_1) \leq \theta \quad \text{and} \quad \frac{\pi}{2} \leq \arg(\rho e^{i\theta} - r_2) \leq \theta.$$

Thus,  $\pi/2 \leq \chi(\rho e^{i\theta}) \leq \theta \leq \pi$ . Moreover since  $r_1, r_2 \geq \rho \cos \theta$  then

$$|\rho e^{i\theta} - r_1| \leq \rho \quad \text{and} \quad |\rho e^{i\theta} - r_2| \leq \rho,$$

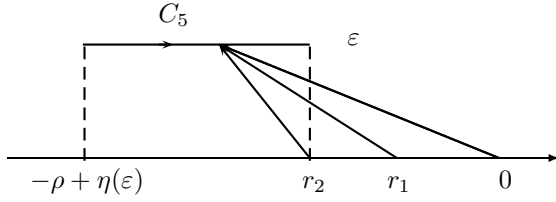
and using the monotony of  $\cos$  on  $[\pi/2, \pi]$ , we have

$$\begin{aligned} \operatorname{Re}(\sqrt{(\rho e^{i\theta} - r_1)(\rho e^{i\theta} - r_2)}) &= |\rho e^{i\theta} - r_1|^{1/2} |\rho e^{i\theta} - r_2|^{1/2} \cos(\chi(\rho e^{i\theta})) \\ &\geq |\rho e^{i\theta} - r_1|^{1/2} |\rho e^{i\theta} - r_2|^{1/2} \cos \theta \\ &\geq \rho \cos \theta. \end{aligned}$$

The above estimation is analogous to (3.12). Now the conclusion follows as in Step 3.

**Step 5.** Here we consider  $C_5(\rho, \varepsilon)$  and  $C_9(\rho, \varepsilon)$ .

Over  $C_5(\rho, \varepsilon)$  we take the parametrization  $s = u + \varepsilon i$ ,  $-\rho + \eta(\varepsilon) \leq u \leq r_2$  :



Notice that

$$\frac{\pi}{2} \leq \arg(u + \varepsilon i - r_1) \leq \pi \quad \text{and} \quad \frac{\pi}{2} \leq \arg(u + \varepsilon i - r_2) \leq \pi,$$

then  $\pi/2 \leq \chi(u + \varepsilon i) \leq \pi$ . Since cosine is negative and decreasing on  $[\pi/2, \pi]$  we have, for  $\varepsilon < |r_1 - r_2|$ ,

$$\begin{aligned} &-a \operatorname{Re}(\sqrt{(u + \varepsilon i - r_1)(u + \varepsilon i - r_2)}) \\ &= a |u + \varepsilon i - r_1|^{1/2} |u + \varepsilon i - r_2|^{1/2} (-\cos(\chi(u + \varepsilon i))) \\ &\leq 2^{1/2} a |u - r_1|. \end{aligned}$$

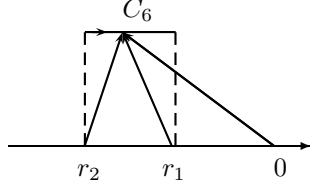
Hence

$$(3.13) \quad \begin{aligned} |e^{ts} h_M(s)| &\leq \frac{e^{(t-a)u - a \operatorname{Re}(\sqrt{(u + \varepsilon i - r_1)(u + \varepsilon i - r_2)})}}{|u + \varepsilon i|} \\ &\leq \frac{e^{(t-a)r_2 + 2^{1/2} a |\rho - r_1|}}{|r_2|}. \end{aligned}$$

By (3.13) we can apply the dominated convergence theorem and since

$$\arg(u + \varepsilon i - r_1) \rightarrow \pi \quad \text{and} \quad \arg(u + \varepsilon i - r_2) \rightarrow \pi, \quad \text{as } \varepsilon \rightarrow 0,$$





Notice that for  $\varepsilon < 1$ ,

$$\begin{aligned}
 |e^{ts} h_M(s)| &= \left| \frac{e^{(t-a)(u+\varepsilon i) - a\sqrt{(u+\varepsilon i - r_1)(u+\varepsilon i - r_2)}}}{u + \varepsilon i} \right| \\
 &= \frac{e^{(t-a)u - a|u+\varepsilon i - r_1|^{1/2}|u+\varepsilon i - r_2|^{1/2} \cos \chi(u+\varepsilon i)}}{|u + \varepsilon i|} \\
 &\leq \frac{e^{(t-a)r_1} e^{-a|u+\varepsilon i - r_1|^{1/2}|u+\varepsilon i - r_2|^{1/2} \cos \chi(u+\varepsilon i)}}{|r_1|} \\
 &\leq \frac{e^{(t-a)r_1} e^{a|u+\varepsilon i - r_1|^{1/2}|u+\varepsilon i - r_2|^{1/2}}}{|r_1|} \\
 (3.14) \quad &\leq \frac{e^{-(t-a)r_1 + a(1+|r_1 - r_2|)^{1/2}}}{|r_1|}.
 \end{aligned}$$

Since

$$\arg(u + \varepsilon i - r_1) \rightarrow \pi \quad \text{and} \quad \arg(u + \varepsilon i - r_2) \rightarrow 0, \quad \text{as } \varepsilon \rightarrow 0,$$

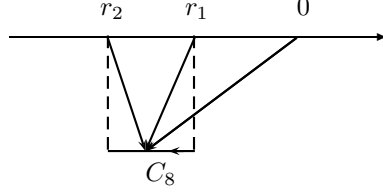
then

$$\begin{aligned}
 \sqrt{(u + \varepsilon i - r_1)(u + \varepsilon i - r_2)} &= |u + \varepsilon i - r_1|^{1/2} |u + \varepsilon i - r_2|^{1/2} \exp(i\chi(u + \varepsilon i)) \\
 &\rightarrow |u - r_1|^{1/2} |u - r_2|^{1/2} i, \quad \text{as } \varepsilon \rightarrow 0.
 \end{aligned}$$

Due to (3.14) we are able to apply the dominated convergence theorem:

$$\begin{aligned}
 \lim_{\varepsilon \rightarrow 0} \int_{C_6(\varepsilon)} e^{ts} h_M(s) ds &= \lim_{\varepsilon \rightarrow 0} \int_{r_2}^{r_1} \frac{e^{(t-a)(u+\varepsilon i) - a\sqrt{(u+\varepsilon i - r_1)(u+\varepsilon i - r_2)}}}{u + \varepsilon i} du \\
 &= \int_{r_2}^{r_1} \frac{e^{(t-a)u - a|u - r_1|^{1/2}|u - r_2|^{1/2} i}}{u} du.
 \end{aligned}$$

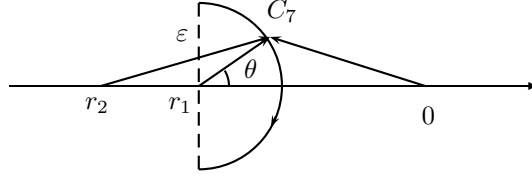
**Step 7.** Now we suppose that  $C_8(\varepsilon)$  is defined by  $s = -u - \varepsilon i$ ,  $-r_1 \leq u \leq -r_2$  :



we can imitate Step 6 to we get

$$\lim_{\varepsilon \rightarrow 0} \int_{C_8(\varepsilon)} e^{ts} h_M(s) ds = - \int_{r_2}^{r_1} \frac{e^{(t-a)u + a|u-r_1|^{1/2}|u-r_2|^{1/2}i}}{u} du.$$

**Step 8.** Finally we deal with  $C_7(\varepsilon)$ . Here, we put  $s = r_1 + \varepsilon e^{-i\theta}$ ,  $-\pi/2 \leq \theta \leq \pi/2$ :



Thus

$$\chi(\varepsilon e^{-i\theta} + r_1) \in \left[ \frac{3}{2}\pi, 2\pi \right] \cup \left[ 0, \frac{1}{2}\pi \right].$$

Consequently, using the fact that  $\cos \theta \geq 0$ , for  $\theta \in [\frac{3}{2}\pi, 2\pi] \cup [0, \frac{1}{2}\pi]$ , we obtain

$$(3.15) \quad \operatorname{Re}(\sqrt{\varepsilon e^{-i\theta}(\varepsilon e^{-i\theta} + r_1 - r_2)}) = \varepsilon^{1/2} |\varepsilon e^{-i\theta} + r_1 - r_2|^{1/2} \cos \chi(\varepsilon e^{-i\theta} + r_1) \geq 0.$$

Also, for  $0 < \varepsilon < -r_1/2$ , we have

$$(3.16) \quad \begin{aligned} |r_1 + \varepsilon e^{-i\theta}| &= \sqrt{r_1^2 + 2r_1\varepsilon \cos(-\theta) + \varepsilon^2} \\ &\geq |r_1 + \varepsilon| = -r_1 - \varepsilon > -\frac{r_1}{2}. \end{aligned}$$

The estimations (3.15), (3.16) and  $t > a$  yield

$$\begin{aligned}
 \left| \int_{C_7(\varepsilon)} e^{ts} h_M(s) ds \right| &\leq \varepsilon e^{(t-a)r_1} \int_{-\frac{\pi}{2}}^{\frac{\pi}{2}} \left| \frac{e^{(t-a)\varepsilon e^{-i\theta} - a\sqrt{\varepsilon e^{-i\theta}(\varepsilon e^{-i\theta} + r_1 - r_2)}}}{r_1 + \varepsilon e^{-i\theta}} \right| d\theta \\
 &\leq \frac{2\varepsilon e^{(t-a)r_1}}{|r_1|} \int_{-\frac{\pi}{2}}^{\frac{\pi}{2}} \left| e^{(t-a)\varepsilon e^{-i\theta} - a\sqrt{\varepsilon e^{-i\theta}(\varepsilon e^{-i\theta} + r_1 - r_2)}} \right| d\theta \\
 &= \frac{2\varepsilon e^{(t-a)r_1}}{|r_1|} \int_{-\frac{\pi}{2}}^{\frac{\pi}{2}} e^{(t-a)\varepsilon \cos(-\theta) - a\operatorname{Re}(\sqrt{\varepsilon e^{-i\theta}(\varepsilon e^{-i\theta} + r_1 - r_2)})} d\theta \\
 &\leq \frac{2\varepsilon e^{(t-a)(r_1 + \varepsilon)}}{|r_1|} \pi.
 \end{aligned}$$

Therefore  $\lim_{\varepsilon \rightarrow 0} \int_{C_7(\varepsilon)} e^{ts} h_M(s) ds = 0$ .

**Step 9.** To finish the proof, we only need to take into account Steps 1-8, together with (3.5), and Propositions 1 and 3.  $\square$

**3.2. Proof of (3.4).** Let us define

$$K(s) = s + b + \sqrt{(s - r_1)(s - r_2)}, \quad s \in \mathbb{C} \setminus [r_2, r_1].$$

Since

$$2b + r_1 + r_2 = 0 \neq -4\lambda rc = r_1 r_2 - b^2,$$

then

$$(s + b)^2 \neq (s - r_1)(s - r_2), \quad \forall s \in \mathbb{C}.$$

This implies that  $1/K$  is analytic over  $\mathbb{C} \setminus [r_2, r_1]$ . Moreover, it is not difficult to see that, for  $\rho$  large enough, we have

$$(3.17) \quad |K(s)| \geq \begin{cases} 2\sqrt{\lambda rc}, & s \in C_2 \cup C_4 \cup C_5 \cup C_7 \cup C_9 \cup C_{10} \cup C_{12}, \\ b, & s \in C_1 \cup C_{13}, \\ (\rho + r_1)^{1/2}(\rho + r_2)^{1/2} \sin \frac{\pi}{4}, & s \in C_3 \cup C_{11}, \\ \varepsilon + |\operatorname{Re}(s) - r_1|^{1/2} |\operatorname{Re}(s) - r_2|^{1/2} \sin \frac{\pi}{4}, & s \in C_6 \cup C_8. \end{cases}$$

As in the proof of (3.3) we have by (3.2)

$$\begin{aligned}
 P(Y^x(t) = 0) &= 1 - \frac{1}{2\pi i} \int_{\sigma - i\infty}^{\sigma + i\infty} e^{ts} \frac{\lambda e^{-(x-x_0)v_s^-}}{cs(v_s^+ + r)} ds \\
 &= 1 - \frac{\lambda e^{-a(\lambda - rc)}}{\pi i} \int_{\sigma - i\infty}^{\sigma + i\infty} e^{ts} h_m(s) ds,
 \end{aligned}$$

where

$$h_m(s) = \frac{e^{-as + a\sqrt{(s-r_1)(s-r_2)}}}{sK(s)}.$$

Finally, the result follows from Proposition 1, (3.17) and the proof of (3.3). Observe that  $a < 0$  and  $t - 2a > 0$ , together with (3.17), allow us to copy, line by line, the proof of (3.3) to show that (3.4) is also true.  $\square$

**Acknowledgement 1.** *The authors would like to thanks Cinvestav-IPN and Universidad Autónoma de Aguascalientes for their hospitality during the realization of this work.*

## REFERENCES

- [1] S. Asmussen (1984). Approximations for the probability of ruin within finite time, *Scand. Act. J.* **1**, 31-57.
- [2] S. Asmussen (2000). *Ruin Probabilities*, World Scientific Publishing Co., Singapore.
- [3] S.N. Chiu, C. Yin (2002). On occupation times for a risk process with reserve-dependent premium, *Stochastic Models* **18**(2), 245-255.
- [4] D.C.M. Dickson (2007). Some finite time ruin problems, *Research Paper Series of Faculty of Economics & Commerce, The University of Melbourne* **153**.
- [5] D.C.M. Dickson, B.D. Hughes, Z. Lianzeng (2005). The density of the time to ruin for a Sparre Andersen process with Erlang arrivals and exponential claims, *Scan. Actuarial J.* **5**, 358-376.
- [6] D.C.M. Dickson, H.R. Waters (1992). The probability and severity of ruin in finite and in infinite time, *Astin Bulletin* **22**(2), 177-190.
- [7] D.C.M. Dickson, G.E. Willmot (2005). The density of the time to ruin in the classical Poisson risk model, *Astin Bulletin* **35**(1), 45-60.
- [8] A.E. dos Reis (1993). How long is the surplus below zero?, *Insurance: Mathematics and Economics* **12**, 23-38.
- [9] S. Drekcic and G.E. Willmot (2003). On the density and moments of the time of ruin with exponential claims, *Astin Bulletin* **33**(1), 11-21.
- [10] H.U. Gerber (1990). When does the surplus reach a given target?, *Insurance: Mathematics and Economics* **9**, 115-119.
- [11] Z.G. Ignatov, V.K. Kaishev (2004). A finite-time ruin probability formula for continuous claim severities, *J. Appl. Prob.* **41**, 570-578.
- [12] H.L. Seal (1972). Numerical calculations of the probability of ruin in the Poisson/Exponential case, *Mitt. Verein Schweiz. Versich. Math.* **72**, 77-100.
- [13] M.R. Spiegel (1988). *Laplace Transforms*, McGraw-Hill.

CINVESTAV-IPN, DEPARTAMENTO DE CONTROL AUTOMÁTICO, APARTADO POSTAL 14-740, MÉXICO D.F., MEXICO

*E-mail address:* `jleon@ctrl.cinvestav.mx`

UNIVERSIDAD AUTÓNOMA DE AGUASCALIENTES, DEPARTAMENTO DE MATEMÁTICAS Y FÍSICA, AV. UNIVERSIDAD 940, C.P. 20100 AGUASCALIENTES, AGS., MEXICO

*E-mail address:* `jvilla@correo.uaa.mx`